

# Bootstrap CIs

## ST552 Lecture 13

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# Today

- ▶ Finish causal inference
- ▶ Bootstrap intervals

# Bootstrap confidence intervals

What if  $\epsilon$  are not from a Normal distribution?

The central limit theorem kicks in, so with large samples, even when the errors aren't Normal,

$$\hat{\beta} \sim N(\beta, \sigma^2 (X^T X)^{-1})$$

The bootstrap is one approach to estimate the sampling distribution of  $\hat{\beta}$ .

# Outline

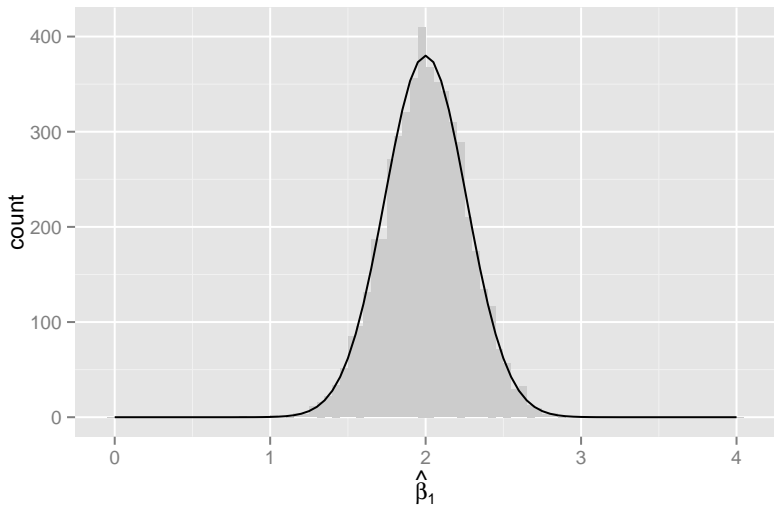
- ▶ What do we do if we know everything? Simulation.
- ▶ How does the bootstrap approximate that process?
- ▶ In practice
- ▶ Limitations

# Simulation

To understand the sampling distribution of  $\hat{\beta}$  we could use simulation.

Just like in HW#4. We know  $\beta$  and the distribution of  $\epsilon$ .

1. Fix  $X$
2. For  $k = 1, \dots, B$ 
  - 2.1 Generate errors,  $\epsilon_i \stackrel{i.i.d}{\sim} \text{Normal}(0, \sigma^2)$
  - 2.2 Construct  $y$ , using the model,  $y = X\beta + \epsilon$
  - 2.3 Use least squares to find  $\hat{\beta}_{(k)}^*$
3. Examine the distribution of  $\hat{\beta}^*$  and compare to  $\beta$ .



```
> quantile(ests$X1, c(0.025, 0.975))  
      2.5%      97.5%  
1.455147 2.539186
```

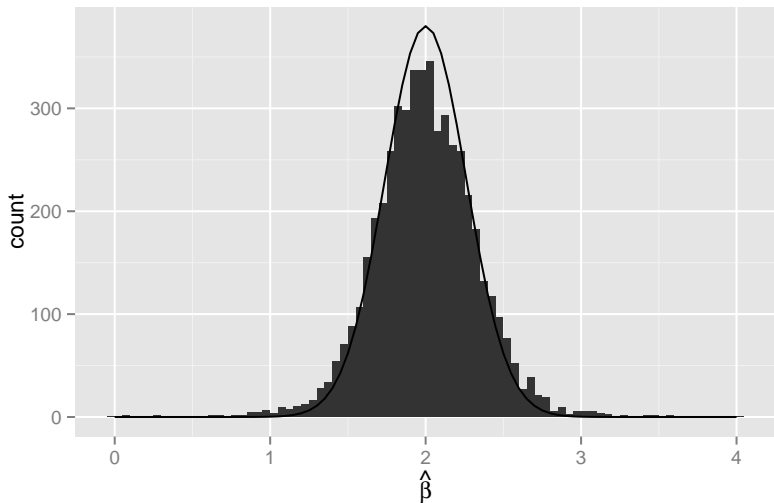
# Simulation

If we want to know what happens to the distribution of  $\hat{\beta}$  when the errors aren't Normal, we could assume some distribution for them and use simulation.

So, swap out step 2.1 for some other distribution. Let's say, Student's t with 3 d.f.

Just like in HW#4. We know  $\beta$  and the distribution of  $\epsilon$ .

1. Fix  $X$
2. For  $k = 1, \dots, B$ 
  - 2.1 Generate errors,  $\epsilon_i \stackrel{i.i.d}{\sim}$  Student's-t<sub>3</sub>
  - 2.2 Construct  $y$ , using the model,  $y = X\beta + \epsilon$
  - 2.3 Use least squares to find  $\hat{\beta}_{(k)}^*$
3. Examine the distribution of  $\hat{\beta}^*$  and compare to  $\beta$



```
> quantile(ests_t$X1, c(0.025, 0.975))  
  2.5%   97.5%  
1.355579 2.631276
```



# Bootstrapping regression

In a real life application we don't know  $\beta$  or the actual distribution of the errors. But we have some reasonable guesses we could make.

0. Fit model and find  $\hat{\beta}$  and  $e_i$

1. Fix  $X$ ,

2. For  $k = 1, \dots, B$

2.1 Generate errors,  $\epsilon_i$  sampled with replacement from  $e_i$

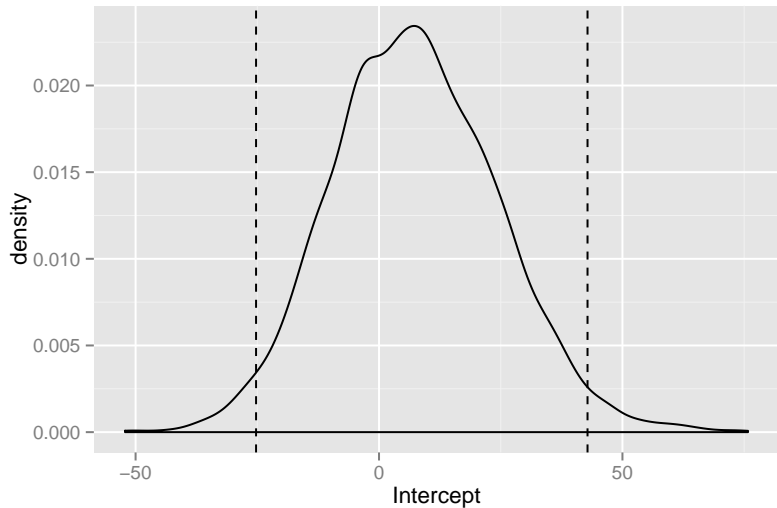
2.2 Construct  $y$ , using the model,  $y = \hat{y} + \epsilon$

2.3 Use least squares to find  $\hat{\beta}_{(k)}^*$

3. Examine the distribution of  $\hat{\beta}^*$  and compare to  $\hat{\beta}$

A naive confidence interval for  $\beta_j$  is the 2.5% and 97.5% quantiles of the distribution of  $\hat{\beta}^*$ . (*This relies on  $E(\hat{\beta}^*) = \hat{\beta}$ , and there are better methods*)

## Example - Faraway



# A reminder of the bootstrap idea

We don't know the distribution of some random variable  $Z$  but we can estimate it with observations of the random variable

$$Z_i, \quad i = 1, \dots, n.$$

Usually, we think about this as using the empirical c.d.f. of  $Z_i$  to approximate the true c.d.f. of  $Z$ .

In practice, sampling from a random variable with a c.d.f. defined as the empirical c.d.f. of a set of numbers,  $Z_i$ , boils down to sampling with replacement from  $Z_i$ .

# Limitations

We might rely on bootstrap confidence intervals when we are worried about the assumption of Normal errors. But, there are limitations.

- ▶ We still rely on the assumption that the errors are independent and identically distributed.
- ▶ Generally scaled residuals are used (residuals don't have the same variance, more later)
- ▶ An alternative bootstrap resamples the  $(y_i, x_{i1}, \dots, x_{ip})$  vectors.